

Exercises

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1.10) We can write

$$A \Delta B = (A \cup B) \cap (\Omega \setminus A \cap B)$$

If A and B are events Then

$A \cup B$ and $A \cap B$ are events. and

So $\Omega \setminus A \cap B$ is an event and

finally

$A \Delta B$ is an event.

1.17) We should check (a), (b) and (c) in the definition.

(a) evident

(b) $P(\emptyset) = 0$ is evident, $P(\Omega) = 1$ follows

(c) Let $A = \bigcup_i A_i$ with $A_i \cap A_j = \emptyset$.

If $\omega \in A$ Then there exists

one and only one A_j such that

$\omega \in A_j$. † Thus

$$Q(A) = \sum_{i: \omega_i \in A} p_i = \sum_j \sum_{i: \omega_i \in A_j} p_i = \sum_j Q(A_j)$$

In particular we never used that

\mathcal{F} is the power set of Ω . For

example we can take

$$\mathcal{F} = \{ \emptyset, \Omega \}$$

and Q would be a probability measure on \mathcal{F} .

1.21) The event needed is

$$D = \left((A \cap B) \cup (A \cap C) \cup (A \cap D) \right) \setminus (A \cap B \cap C)$$

$$= D_1 \cup D_2 \cup D_3 \quad \text{where}$$

$$D_1 = (A \cap B) \setminus (A \cap B \cap C)$$

$$D_2 = (A \cap C) \setminus (A \cap B \cap C)$$

$$D_3 = (A \cap D) \setminus (A \cap B \cap C)$$

Observe that $D_i \cap D_j = \emptyset$ if $i \neq j$.

Moreover

$$P(D_1) = P(A \cap B) - P(A \cap B \cap C)$$

since $A \cap B \cap C \subset A \cap B$. Similarly

$$P(D_2) = P(A \cap C) - P(A \cap B \cap C)$$

$$P(D_3) = P(B \cap C) - P(A \cap B \cap C)$$

so that

$$P(D) = P(A \cap B) + P(B \cap C) + P(A \cap C) -$$

$$3 \cdot P(A \cap B \cap C) = \frac{6}{10} - \frac{3}{5}$$

1.27) The Total number of possible bridge hands is

$$\frac{52!}{13!^4}$$

On The other hand There are

$4! = 24$ way To give an ace to

④

each player and

$$\frac{48!}{12!^4}$$

way to distribute the remaining 48 cards to the 4 players, 12 each.

Thus the probability is

$$\frac{4! \frac{48!}{12!^4}}{\frac{52!}{13!^4}} = \frac{24 \cdot 13^4 \cdot 48!}{52!}$$

1.30)

(a) The probability is

$$1 - \left(1 - \frac{1}{6}\right)^4 = 1 - \left(\frac{5}{6}\right)^4 = \approx 518$$

(b) The probability is

$$1 - \left(1 - \frac{1}{36}\right)^{24} = 1 - \left(\frac{35}{36}\right)^{24} = 0.491$$

So ~~a~~ a is greater than ~~b~~ b.

Observe that for small x we

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have

$$(1-x)^n \approx 1 - nx + O(x^2)$$

Thus at this level of approximation
a and b are equal. In other words

The probability of exactly 1 six in
for throws equal the probability of
exactly one double six in 24 double
throws.

2.46) We have

$$(A \cap (\Omega \setminus B)) \cup (A \cap B) = A$$

wh. i.e.

$$(A \cap (\Omega \setminus B)) \cap (A \cap B) = \phi$$

So we have

$$P(A \cap (\Omega \setminus B)) + P(A \cap B) = P(A)$$

Thus if $P(A \cap B) = P(A)P(B)$ Then

we have

$$\begin{aligned}
P(A \cap (\Omega \setminus B)) &= P(A) - P(A)P(B) = \\
&= P(A)(1 - P(B)) = \\
&= P(A)P(\Omega \setminus B)
\end{aligned}$$

since $P(\Omega \setminus B) = 1 - P(B)$.

1.52(a) ~~best~~ Define the events

$$B_1 = \{ \text{first pick is Black} \}$$

$$W_1 = \{ \text{first pick is White} \}$$

$$B_2 = \{ \text{second pick is Black} \}$$

$$W_2 = \{ \text{second pick is White} \}$$

Thus.

$$P(B_2) = P(B_2 | B_1) P(B_1) +$$

$$P(B_2 | W_1) P(W_1) =$$

$$= \frac{2}{9} \cdot \frac{4}{7} + \frac{6}{9} \cdot \frac{3}{7} = \frac{46}{63}$$

(b) Define

$$A_I = \{ \text{you also pick urn I} \}$$

$$A_{II} = \{ \text{you pick urn II} \}$$

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and

$B = \{ \text{you pick Black} \}$

$W = \{ \text{you pick White} \}$

$$P(A_I | B) = \frac{P(B | A_I) \cdot P(A_I)}{P(B | A_I) \cdot P(A_I) + P(B | A_{II}) \cdot P(A_{II})}$$

$$= \frac{\frac{4}{7} \cdot \frac{1}{2}}{\frac{4}{7} \cdot \frac{1}{2} + \frac{6}{8} \cdot \frac{1}{2}} = \frac{16}{37}$$

Problems

9. Comments on The solution in
The book:

a) \bar{P}_k If the coin is fair The prob.
of getting k head is equal To
The prob of getting k Tail,
That is $n-k$ head. Thus

$$P(Y=k) = P(Y=n-k)$$

b) If X and Y are 2 r.v we have

$$\begin{aligned}
 P(X+Y=k) &= \sum_{\substack{x,y \\ x+y=k}} P(X=x)P(Y=y) \\
 &= \sum_x P(X=x)P(Y=k-x)
 \end{aligned}$$

14) a) Assume The formula is True

for $n \leq N$ we will prove it for $n = N$.

We can rewrite The equation as

$$P\left(\bigcup_{i=1}^n A_i\right) = \sum_{I \subset \{1, \dots, n\}} (-1)^{|I|-1} P\left(\bigcap_{i \in I} A_i\right)$$

where The sum is over all subset of $\{1, 2, \dots, n\}$ and $|I|$ is The cardinality of I (i.e. The number of elements in I).

$$P\left(\bigcup_{i=1}^N A_i\right) = P\left(\bigcup_{i=1}^{N-1} A_i\right) + P(A_N) - P\left(\left(\bigcup_{i=1}^{N-1} A_i\right) \cap A_N\right) \quad (9)$$

but

$$\left(\bigcup_{i=1}^{N-1} A_i\right) \cap A_N = \bigcup_{i=1}^{N-1} (A_i \cap A_N)$$

so that using The inductive

assumption we get

$$P\left(\bigcup_{i=1}^N A_i\right) = P(A_N) + \sum_{I \subseteq \{1, \dots, N-1\}} (-1)^{|I|-1} P\left(\bigcap_{i \in I} A_i\right)$$

$$- \sum_{I \subseteq \{1, \dots, N-1\}} (-1)^{|I|-1} P\left(\bigcap_{i \in I} (A_i \cap A_N)\right)$$

but

$$\bigcap_{i \in I} (A_i \cap A_N) = \left(\bigcap_{i \in I} A_i\right) \cap A_N$$

Thus The last sum becomes

$$- \sum_{I \subseteq \{1, \dots, N-1\}} (-1)^{|I|-1} P\left(\bigcap_{i \in I} (A_i \cap A_N)\right) =$$

$$\sum_{\substack{J \subseteq \{1, \dots, N\} \\ N \in J}} (-1)^{|J|-1} P\left(\bigcap_{j \in J} A_j\right)$$

The statement easily follows.

b) Let $A_i = \{i \text{th Key is an its hook}\}$

we need

$$P(\bigcup_{i=1}^n A_i)$$

but $P(\bigcap_{i \in I} A_i) = \left(\frac{1}{n}\right)^{|I|}$ so that

$$P(\bigcup_{i=1}^n A_i) = \sum_{k=1}^n (-1)^{k-1} \binom{n}{k} \frac{1}{n^k} =$$

$$1 - \left(1 - \frac{1}{n}\right)^n = 1 - P(\bigcap_{i=1}^n A_i)$$

Clearly

$$\lim_{n \rightarrow \infty} P(\bigcup_{i=1}^n A_i) = 1 - e^{-1}$$

The following morning we have the same situation but this time

$$P\left(\bigcap_{i \in I} A_i\right) = \frac{(n - |I|)!}{i!}$$

since no two keys can be on the same hook. Thus

$$\begin{aligned} P\left(\bigcup_i A_i\right) &= \sum_{k=1}^n (-1)^{k-1} \binom{n}{k} \frac{(n-k)!}{n!} = \\ &= 1 - \sum_{k=0}^n (-1)^k \frac{1}{k!} \end{aligned}$$

So that again

$$\lim_{n \rightarrow \infty} P\left(\bigcup_i A_i\right) = 1 - e^{-1}.$$

Since the probability that no key was hung on its own hook is

$$P\left(\bigcap_i A_i^c\right) = 1 - P\left(\bigcup_i A_i\right) = \sum_{k=0}^n (-1)^k \frac{1}{k!}$$

so that

$$\lim_{n \rightarrow \infty} P\left(\bigcap_i A_i^c\right) = e^{-1}$$

17) If the first Toss is head there are 2 possibilities:

- 1) It is followed by $r-1$ heads
- 2) There is at least one tail in the following $r-1$ Tosses

In the first case we have a success.

In the second case we can start again after we see the first Tail.

Observe that, given that there is at least one Tail in the first $r-1$ Tosses, the Tosses after the first Tail are independent. Thus we

get

$$P(E | A = \text{head}) = p^{r-1} + (1-p^{r-1})P(E | A = \text{Tail})$$

On the other hand if The first 12
 toss is tail, there need to be
 at least another a head in the
 following $s-1$ Tosses so that

$$P(E | A = \text{tail}) = (1 - q^{s-1}) P(E | A = \text{head})$$

Thus

$$P(E) = P(E | A = \text{tail}) q + P(E | A = \text{head}) p$$

but

$$P(E | A = \text{head}) = p^{r-1} + (1 - p^{r-1})(1 - q^{s-1}) P(E | A = \text{head})$$

so that

$$P(E | A = \text{head}) = \frac{p^{r-1}}{1 - (1 - p^{r-1})(1 - q^{s-1})}$$

$$P(E | A = \text{tail}) = \frac{(1 - q)^{s-1} p^{r-1}}{1 - (1 - p^{r-1})(1 - q^{s-1})}$$

$$P(E) = \frac{(p^r + q(1 - q^{s-1})p^{s-1})}{(p^{r-1} + q^{s-1} - p^{r-1}q^{s-1})}$$