Transportation Distance and the Central Limit Theorem

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Abstract

For probability measures on a complete separable metric space, we present sufficient conditions for the existence of a solution to the Kantorovich transportation problem. We also obtain sufficient conditions (which sometimes also become necessary) for the convergence of probability measures in the transportation distance when the cost function is a continuous and non-decreasing function of the metric. As an application, the CLT in the transportation distance is proved for independent, strongly mixing and associated sequences.

Keywords: Kantorovich transportation problem, convergence in transportation distance, Central Limit Theorem in transportation distance, Wasserstein distance, strong mixing sequences, associated sequences.

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1 Introduction

Let (M, d) be a metric space and let $c : M \times M \to \mathbf{R}$, be a non-negative Borel function. The transportation c-distance $T_c(\mu, \nu)$ between two probability measures μ and ν defined on the Borel σ-field $\mathcal{B}(M)$ is given via

$$T_c(\mu, \nu) = \inf \mathbf{E}c(X, Y).$$

Above, the infimum is taken over all M-valued random elements X and Y defined on the probability space $(\Omega, \mathcal{F}, \mathbf{P})$ and having, respectively, μ and ν for probability distribution. In other words,

$$T_c(\mu, \nu) = \inf_{\Pi} \int c(x, y) d\pi(x, y), \tag{1}$$

where the infimum is taken over the set Π of all probability measures on $\mathcal{B}(M \times M)$ with marginals μ and ν. The transportation distance is related to the celebrated Kantorovich transportation problem: if μ and ν are two distributions of mass and if $c(x, y)$ represents the cost of transporting a unit of mass from the location x to the location y, what is the minimal

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total transportation cost to transfer μ to ν? The minimal total transportation cost is exactly the transportation distance corresponding to the cost function c.

The c-transportation distance with $c(x,y) = d^p(x,y)$, $p \geq 1$, is associated to the Wasserstein or Mallows p-distance W_p, $W_p(\mu,\nu) = (T_{d^p}(\mu,\nu))^{1/p}$. If M is the real line $\mathbf{R}$ with the Euclidean distance, the Wasserstein-Mallows p-distance between two distribution functions F and G has the following useful representation

$$W_p^p(F,G) = \int_0^1 |F^{-1}(t) - G^{-1}(t)|^p dt, \tag{2}$$

where the inverse transformation of F is defined as

$$F^{-1}(t) = \sup\{x \in \mathbf{R} : F(x) \leq t\}.$$

The representation (2) was obtained when $p = 1$ by Salvemini [20] (for discrete distributions) and by Dall'Aglio [7] (in the general case), while for $p = 2$ it is due to Mallows [14]. It implies that the random variables $X = F^{-1}(U)$ and $Y = G^{-1}(U)$, where U is a uniform random variable on $(0,1)$, are minimizers of the total transportation cost in the transportation problem. Major [13] generalized (2) to a convex cost function $c(x,y) = c(x-y)$:

$$T_c(F,G) = \int_0^1 c(F^{-1}(t) - G^{-1}(t))dt.$$

The representation (2) is an important tool in proving the following convergence result.

Let $p = 1, 2$ and let F_n, F be distribution functions on $\mathbf{R}$ such that for any n, $\int |x|^p dF_n < +\infty$, and $\int |x|^p dF < +\infty$. Then

$$W_p(F_n, F) \to 0 \iff \begin{cases} (a)\ F_n \Longrightarrow F, \\ (b) \int |x|^p dF_n \to \int |x|^p dF. \end{cases} \tag{3}$$

For $p = 1$ the equivalence (3) was proved by Dobrushin [8], while for $p = 2$ it is due to Mallows [14].

Bickel and Freedman [2] extended the statement (3) to probability measures μ_n and μ defined on a separable Banach space $(\mathbf{B}, \|\cdot\|)$ and to all $p \in [1, +\infty)$ as follows:

Let $1 \leq p < \infty$, and let $\int \|x\|^p \mu_n(dx) < \infty$, $\int \|x\|^p \mu(dx) < \infty$. Then $W_p(\mu_n, \mu) \to 0$ as $n \to \infty$ is equivalent to each of the following.

(a) $\mu_n \Longrightarrow \mu$ and $\int \|x\|^p \mu_n(dx) \to \int \|x\|^p \mu(dx)$.

(b) $\mu_n \Longrightarrow \mu$ and $\|\cdot\|^p$ is uniformly μ_n-integrable.

(c) $\int \phi(x)\mu_n(dx) \to \int \phi(x)\mu(dx)$ for every continuous ϕ such that $\phi(x) = O(\|x\|^p)$ at infinity.

Since in general an analog of the representation (2) does not exist for probability measures on a Banach space, Bickel and Freedman proved, in their setting, the existence of a solution to the transportation problem for $c(x,y) = \|x-y\|^p$.

More recently, Ambrosio, Gigli, and Savaré proved ([1], Proposition 7.1.5) an analog of part (b) of the above result for probability measures on a Radon space X (see also Lemma 5.1.7 and Remark 7.1.11 there). These authors also established the existence of a solution to the Kantorovich transportation problem in X for a wide class of cost functions. We use this existence result to prove criteria for the convergence in T_c with $c(x,y) = C(d(x,y))$, where C is a non-decreasing continuous function satisfying the doubling condition (6) which controls the rate of growth of C (Theorem 2 and Corollary 1). Since the class of such cost functions includes all the d^ps, $p \geq 1$, the convergence results of Bickel and Freedman as well as those of Ambrosio *et al.* follow from our Corollary 1. Note that instead of the Radon space X (a separable metric space where, by definition, every probability measure is tight), we consider here a more familiar probabilistic object, a complete separable space (M,d). In our framework, completeness and separability together will provide tightness; all our arguments remain true for Radon spaces (see also Remark 1).

In Theorem 2 we also obtain sufficient conditions for the convergence of probability measures in the transportation distance without assuming the doubling condition on C. For instance, any convex $C : \mathbf{R}^+ \to \mathbf{R}^+$ with $C(0) = 0$ satisfies Theorem 2. We then provide an example of a C growing exponentially fast for which the converse implication does not hold.

2 Convergence in Transportation Distance

The following result by Ambrosio *et al.* [1] asserts the existence in Π of a probability measure which minimizes the total transportation cost under rather weak assumptions on the cost function. For the sake of completeness, we include a self-contained proof in Section 4.

Theorem 1. *Let (M,d) be a complete separable metric space, and let $T_c(\mu,\nu)$ be defined by (1) with $c : M \times M \to [0,+\infty)$ lower semicontinuous. Then, there exists $\pi^* \in \Pi$ such that $\int c(x,y)d\pi^*(x,y) = T_c(\mu,\nu)$. Equivalently, there exists a pair of random elements X and Y with respective distribution μ and ν, such that $\mathbf{E}c(X,Y) = T_c(\mu,\nu)$.*

Remark 1. In the corresponding statement in [1] the space (M,d) need not be a complete separable metric space but just a Radon space. In fact, our proof also shows that completeness is unnecessary and that the tightness of μ and ν will suffice. On the other hand, the hypothesis of separability of (M,d) can be weakened to topological separability if both μ and ν have separable supports (see Billingsley [3], Appendix III).

The Kantorovich problem is closely related to the Monge transportation problem which is the problem of finding a map s^* pushing μ forward to ν (i.e. such that $\nu(B) = \mu(s^{-1}(B))$ for any Borel set B) and minimizing the total transportation cost: $\inf_s \int c(x,s(x))d\mu = \int c(x,s^*(x))d\mu$, where the infimum is taken over all Borel maps s pushing μ forward to ν. A solution s^* to the Monge transportation problem uniquely determines a probability measure π^* on $M \times M$ such that the random elements X and Y, $Y = s^*(X)$, with respective distributions μ and ν have joint law π^*. This measure π^* minimizes the Monge transportation cost:

$$\int c(x,y)d\pi^*(x,y) = \inf_{\Pi^*} \int c(x,y)d\pi(x,y), \tag{4}$$

where the infimum is taken over the set Π^* of joint distributions of M-valued random elements X and Y with respective distributions μ and ν and such that Y is measurable with respect to the Borel field $\sigma(X)$. Comparing (4) and (1) yields the relation $\Pi^* \subset \Pi$, which immediately leads to the following conclusions: (i) the (Kantorovich) transportation distance $T_c(\mu,\nu)$

never exceeds the Monge transportation distance $\tilde{T}_c(\mu,\nu)$,

$$\tilde{T}_c(\mu,\nu) = \inf_{\Pi^*} \int c(x,y)d\pi(x,y) = \inf_s \int c(x,s(x))d\mu;$$

(ii) a probability measure π^* corresponding to the solution s^* of the Monge transportation problem (MTP) is not necessarily a solution to the Kantorovich transportation problem (KTP); conversely, a solution π' of the KTP, where π' is the joint distribution of X and Y, is a solution to the MTP if and only if there exists a Borel map s' such that $Y = s'(X)$.

For random elements X and Y in a Hilbert space Cuesta and Matran [6] have given conditions for the existence of an increasing map s, $s(X) = Y$, such that $W_2^2(\mu,\nu) = \mathbf{E}\|X - s(X)\|^2$, i.e. X and $Y = s(X)$ give the solution to both the MTP and the KTP. They also showed that if μ is either absolutely continuous with respect to the Lebesgue measure on $\mathbf{R}^k$ or is a Gaussian measure on a Hilbert space, then these conditions are satisfied. For compactly supported absolutely continuous distributions on $\mathbf{R}^k$ and a convex cost function $c(x-y)$ Caffarelli [5] has determined the form of the optimal map (the solution to the MTP) as a gradient of c; the uniqueness of the solution is also obtained there. Concurrently, Gangbo and McCann [10] proved the same results for non-necessarily boundedly supported probability measures. They also showed that the solution to the MTP is the KTP solution as well, and that a similar result holds true for $c(x,y) = \ell(\|x-y\|)$, where ℓ is strictly concave. Note that in all the existence statements mentioned above, the conditions of Theorem 1 are satisfied. A comprehensive review of the results on the solutions to the KTP and the MTP can be found in the books of Rachev and Rüschendorf [18].

The main result of the work presented here is now given.

Theorem 2. *Let μ_n and μ be probability measures on a complete separable metric space (M,d) and let $c : M \times M \to \mathbf{R}$ be such that $c(x,y) = C(d(x,y))$, where $C : [0,+\infty) \to [0,+\infty)$ is a non-decreasing continuous function with $C(0) = 0$. Let*

$$\int C(2d(x,a))\mu_n(dx) < \infty, \qquad \int C(2d(x,a))\mu(dx) < \infty, \tag{5}$$

for some (and, therefore, for all) $a \in M$. Then

$$\left.\begin{array}{l} (a)\ \mu_n \Longrightarrow \mu, \\ (b) \displaystyle\int C(2d(x,a))\mu_n(dx) \to \int C(2d(x,a))\mu(dx) \end{array}\right\} \Longrightarrow T_c(\mu_n,\mu) \to 0.$$

Conversely, if $T_c(\mu_n,\mu) \to 0$, then $\mu_n \Longrightarrow \mu$. If, additionally, C satisfies a doubling condition, i.e. if there exists a positive constant λ such that for all $y \geq 0$

$$C(2y) \leq \lambda C(y), \tag{6}$$

then

$$T_c(\mu_n,\mu) \to 0 \Longleftrightarrow \begin{cases} (a)\ \mu_n \Longrightarrow \mu, \\ (b)\ \int C(2d(x,a))\mu_n(dx) \to \int C(2d(x,a))\mu(dx). \end{cases}$$

Corollary 1. *If, in the setting of Theorem 2, C satisfies (6), then*

$$(6) \Longleftrightarrow \int C(d(x,a))\mu_n(dx) < \infty, \int C(d(x,a))\mu(dx) < \infty,$$

and thus

$$T_c(\mu_n,\mu) \to 0 \Longleftrightarrow \begin{cases} (a)\ \mu_n \Longrightarrow \mu, \\ (b')\ \int C(d(x,a))\mu_n(dx) \to \int C(d(x,a))\mu(dx). \end{cases}$$

Corollary 1 is equivalent to a result of Rachev (Theorem 1 in [17]) proved by using the relations between the Lévy-Prokhorov metric and the T_c-distance. Since for any $p \geq 1$, the function $c(x,y) = d^p(x,y)$ satisfies the conditions of Theorem 2 as well as (6) with $\lambda = 2^p$, Corollary 1 recovers part (a) in the result of Bickel and Freedman mentioned above. Ambrosio *et al.* proved an analog of Theorem 2 in a Hilbert space when the cost function is continuous, strictly increasing and surjective ([1], Theorem 5.1.13).

Note that the class of functions C covered by Theorem 2 includes functions with an exponential rate of growth at infinity (e.g. $C(d(x,y)) = \exp(d(x,y)) - 1$). For functions C growing exponentially fast at infinity, and in contrast to $C(d(x,y)) = d^p(x,y)$, $T_c(\mu_n,\mu) \to 0$ need not imply the convergence of $\int C(2d(x,a))\mu_n(dx)$ to $\int C(2d(x,a))\mu(dx)$, for some $a \in M$. Indeed, one can take the probability measures μ_n and μ on $\mathbf{R}$ defined in Example 1, below, and $c(x,y) = C(|x-y|) = \exp(|x-y|) - 1$.

As a corollary to Theorem 2 we obtain the following result relating convergence in total variation to convergence in transportation distance. As well known, the total variation distance itself is a particular case of transportation distance (with $c(x,y) = 2\mathbf{1}_{\{x\neq y\}}$).

Corollary 2. *Let μ and ν be compactly supported probability measures on a complete separable metric space (M,d), and let ϕ be a continuous function on (M,d). Then*

$$\left| \int \phi(x)\mu(dx) - \int \phi(x)\nu(dx) \right| \leq L_\phi \|\mu - \nu\|_{TV},$$

for some positive constant L_ϕ.

Let μ_n be probability measures on M with respective compact supports K_n, $n \geq 1$. Let $\cup_n K_n$ be bounded. If $c(x,y) = C(d(x,y))$, where $C : [0,+\infty) \to [0,+\infty)$ is non-decreasing, continuous with $C(0) = 0$, then

$$\|\mu_n - \mu\|_{TV} \to 0 \Rightarrow T_c(\mu_n,\mu) \to 0.$$

Without the boundedness restriction on $\cup K_n$ the last implication is not true, as the following example shows.

Example 1. Let μ be the uniform distribution on $[0,1]$ and, for all $n \in \mathbf{N}$, let

$$\mu_n(dx) = \frac{n-1}{n}\mathbf{1}_{[0,1]}(x)dx + \frac{1}{n}\delta_{x_n}(dx),$$

$x_n \notin [0,1]$. Then

$$\|\mu_n - \mu\|_{TV} = \int_0^1 \frac{1}{n}dx + \mu_n(x_n) = \frac{2}{n} \to 0$$

as $n \to \infty$. Hence $\mu_n \xrightarrow{TV} \mu$, for any choice of the sequence (x_n). Let $c(x,y) = C(|x-y|)$, with $C : [0,+\infty) \to [0,+\infty)$, $C(0) = 0$, convex, also satisfying (6). Then,

$$\int C(|x|)\mu(dx) = \int_0^1 C(|x|)dx \le \max_{0\le|x|\le 1} C(|x|) < +\infty,$$

$$\int C(|x|)\mu_n(dx) = \int_0^1 \frac{n-1}{n}C(|x|)dx + \mu_n(x_n)C(|x_n|) \le \max_{0\le|x|\le 1} C(|x|)\frac{n-1}{n} + \frac{C(|x_n|)}{n} < +\infty,$$

for any n. So all the conditions of Corollary 1 are satisfied. Since weak convergence is implied by convergence in total variation, $T_c(\mu_n,\mu) \to 0$ holds if and only if $\int C(|x|)\mu_n(dx) \to \int C(|x|)\mu(dx)$. Take $x_n = 2^n$, then $C(|x_n|) = C(2^n) \ge 2^{n-1}C(2)$ and $C(|x_n|)/n \to +\infty$ as $n \to \infty$. Therefore,

$$\int C(|x|)\mu_n(dx) \ge \frac{C(|x_n|)}{n} \to +\infty \ne \int C(|x|)\mu(dx).$$

By Corollary 1, $T_c(\mu_n,\mu)$ does not converge to 0.

3 Applications to the Central Limit Theorem

We now apply Theorem 2 to derive the CLT in the transportation distance. We obtain sufficient conditions for the convergence of the distributions functions of the normalized sums to the standard Gaussian distribution in $\mathbf{R}$ for strictly stationary independent, strongly mixing or associated sequences. The CLT is also proved for non-stationary sequences of independent random variables satisfying the Lyapunov condition.

3.1 Independent sequences

3.1.1 Stationary case

Let (X_n) be a sequence of independent identically distributed random variables, $\mathbf{E}X_1 = 0$, $\mathbf{E}X_1^2 = \sigma^2$, $0 < \sigma < +\infty$, $S_n = \sum_{i=1}^n X_i$. Let F_n denote a distribution function of the normalized sum $S_n/(\sigma\sqrt{n})$, and let Φ be a distribution function of $Z \sim N(0,1)$. Theorem 3 below provides additional conditions on the sequence (X_n) and on the cost function to obtain the convergence of F_n to Φ in the transportation distance.

Theorem 3. (i) *If* $\mathbf{E}|X_1|^p < +\infty$ for some $p \ge 2$, then $W_p(F_n,\Phi) \to 0$.
(ii) *Additionally, let* $c(x,y) = C(|x-y|)$*, where* $C : [0,+\infty) \to [0,+\infty)$ *is a non-decreasing continuous function with* $C(0) = 0$ *and* $C(x) = O(x^p)$ *at infinity. Then* $T_c(F_n,\Phi) \to 0$.

The CLT in the W_2-distance was proved by Tanaka [24] for distributions on $\mathbf{R}$ and by Cuesta and Matran [6] for distributions on a Hilbert space; both results require the finiteness of the fourth moment. Recently, Johnson and Samworth [11], [12] proved Theorem 3(i) by using the subadditivity of the Wasserstein distance. We prove this result by using an inequality of Sakhanenko [19]; the arguments of the proof will be generalized to non-stationary sequences in Theorem 4.

3.1.2 Non-stationary case

Let (X_n) be a sequence of centered independent random variables with $0 < \mathbf{E}X_n^2 < +\infty$, for all n, and let F_n be a distribution function of S_n/σ_n, where $\sigma_n^2 = \sum_{i=1}^n \mathbf{E}X_i^2$.

Theorem 4. (i) *Let the random variables X_n, $n \geq 1$, be such that $\mathbf{E}|X_n|^p < +\infty$, for some $p > 2$. Assume, moreover, that the Lyapunov condition,*

$$\lim_{n\to+\infty} \frac{\sum_{i=1}^n \mathbf{E}|X_i|^p}{\sigma_n^p} = 0, \tag{7}$$

holds true. Then $W_p(F_n, \Phi) \to 0$.
(ii) *If, additionally, the function c is as in Theorem* 3 (ii), *then $T_c(F_n, \Phi) \to 0$.*

3.2 Strongly mixing sequences

Let α_n, $n \geq 1$, be the coefficients of strong mixing of the sequence (X_n), i.e.

$$\alpha_n = \sup\{|P(A \cap B) - P(A)P(B)| : A \in \mathcal{F}_1^k,\ B \in \mathcal{F}_{k+n}^{+\infty}, k \geq 1\},$$

where $\mathcal{F}_k^{k+m}$ is the σ-field generated by the random variables $X_k, X_{k+1}, ..., X_{k+m}$. Recall that (X_n) is said to satisfy a strong mixing condition if $\alpha_n \to 0$ as $n \to +\infty$.

Among a variety of versions of the central limit theorem for strong mixing sequences (see the recent survey of Merlevède *et al.* [15] and the references therein) the most suitable for our purpose turns out to be the one proved by Doukhan *et al.* [9]. It takes the following form. Let (X_n) be a strictly stationary strongly mixing sequence with $\mathbf{E}X_1 = 0$, $0 < \mathbf{E}X_1^2 < +\infty$ and such that for a quantile function Q of $|X_1|$, $Q(u) = \inf\{t : P(|X_1| > t) \leq u\}$, the condition

$$\sum_{n=1}^{\infty} \int_0^{\alpha_n} Q^2(u)du < +\infty \tag{8}$$

is satisfied. Then, $\mathbf{E}X_1^2 + 2\sum_{n=2}^{\infty} \mathrm{Cov}(X_1, X_n) = \sigma^2 < +\infty$, $(\mathbf{E}S_n^2)/n \to \sigma^2$, and

$$\frac{S_n}{\sigma\sqrt{n}} \xrightarrow{d} Z \sim N(0,1),$$

where once again $S_n = \sum_{i=1}^n X_i$.

To obtain the convergence of the distribution functions F_n of $S_n/(\sigma\sqrt{n})$ to Φ in the transportation distance, we need an additional restriction on the rate of decay of the mixing coefficient α_n in order to provide a moment bound for sums. Such a result, due to Shao and Yu [22], asserts that if (X_n) is a centered strongly mixing sequence with

$$\mathbf{E}|X_n|^{p+\delta} < +\infty, \quad \alpha_n \leq Cn^{-\frac{p(p+\delta)}{2\delta}}, \tag{9}$$

for some $p > 2$, $\delta > 0$, $n \geq 1$, then

$$\mathbf{E}|S_n|^p \leq Kn^{\frac{p}{2}} \max_{1\leq i\leq n} (\mathbf{E}|X_i|^{p+\delta})^{\frac{p}{p+\delta}}, \tag{10}$$

where the positive constant K depends on p, δ and C only.

Theorem 5. (i) *For any centered strictly stationary random sequence* (X_n) *with strong mixing coefficients satisfying the condition* (9), *we have* $W_p(F_n, \Phi) \to 0$.
(ii) *If, moreover, c is as in Theorem* 3 (ii), *then* $T_c(F_n, \Phi) \to 0$.

To prove Theorem 5, we apply Theorem 2. Since the CLT of Doukhan *et al.* implies the weak convergence to the standard Gaussian distribution, it is sufficient to verify the convergence of the absolute moments of order p for the normalized sums.

3.3 Associated sequences

3.3.1 Negatively associated sequences

Recall that a set of random variables $\xi = (\xi_1, ..., \xi_m)$ is called negatively associated (NA) if for any two coordinatewise increasing functions $f,\ g : \mathbf{R}^m \to \mathbf{R}$ such that $\mathbf{E}f(\xi)g(\xi)$, $\mathbf{E}f(\xi)$ and $\mathbf{E}g(\xi)$ exist,

$$\mathrm{Cov}\,(f(\xi), g(\xi)) \le 0. \tag{11}$$

An infinite set of random variables is NA if all of its finite subsets are NA.

Su *et al.* [23] proved a CLT for NA sequence in the following form. Let (X_n) be a strictly stationary NA sequence with $\mathbf{E}X_1 = 0$ and $0 < \mathbf{E}X_1^2 < +\infty$. Let $\sigma^2 = \mathbf{E}X_1^2 + 2\sum_{n=2}^{\infty} \mathrm{Cov}(X_1, X_n) > 0$ ($\sigma^2 \le \mathbf{E}X_1^2$ follows from (11)). Then $(\mathbf{E}S_n^2)/n \to \sigma^2$, and

$$\frac{S_n}{\sigma\sqrt{n}} \xrightarrow{d} Z \sim N(0,1).$$

For NA sequence (X_n) with $\mathbf{E}|X_n|^p < +\infty$, $n \ge 1$ and $p \ge 1$, Shao [21] established the following property. Let (X_n^*) be a sequence of independent random variables such that, for all $n \ge 1$, X_n^* and X_n have the same distribution. Then

$$\mathbf{E}|S_n|^p \le \mathbf{E}|S_n^*|^p, \tag{12}$$

where $S_n^* = \sum_{i=1}^n X_i^*$.

Theorem 6. *Let* (X_n) *be a centered strictly stationary NA sequence with* $\sigma^2 > 0$ *and* $\mathbf{E}|X_1|^p < +\infty$, $p \ge 2$. *Then, the sequence of distribution functions* F_n *of the normalized sums* $S_n/(\sigma\sqrt{n})$ *converges to* Φ *in* W_p*-distance.*
(ii) *If, additionally, c is as in Theorem* 3 (ii), *then* $T_c(F_n, \Phi) \to 0$.

3.3.2 Positively associated sequences

A finite set of random variables $\xi = (\xi_1, ..., \xi_m)$ is called positively associated (PA) if the inequality (11) holds true with "$\le$" replaced by "$\ge$". An infinite set of random variables is PA if all of its finite subsets are PA.

Newman and Wright [16] obtained the following version of the CLT. Let (X_n) be a centered strictly stationary PA sequence such that $0 < \mathbf{E}X_1^2 < +\infty$ and $\sigma^2 = \mathbf{E}X_1^2 + 2\sum_{n=2}^{\infty} \mathrm{Cov}(X_1, X_n) < +\infty$. Then $(\mathbf{E}S_n^2)/n \to \sigma^2$, and

$$\frac{S_n}{\sigma\sqrt{n}} \xrightarrow{d} Z \sim N(0,1).$$

Asymptotic independence for PA sequence (X_n) is usually stated in terms of the Cox-Grimmett coefficient $u(n)$ defined by:

$$u(n) = \sup_{k\geq 1} \sum_{j:|j-k|\geq n} \mathrm{Cov}(X_j, X_k).$$

For a stationary sequence the Cox-Grimmett coefficient is just the tail of the series of covariances, $u(n) = 2\sum_{k=n+1}^{\infty} \mathrm{Cov}(X_1, X_k)$.

To prove the convergence of F_n to Φ in the transportation distance, we use a condition on the rate of decay of the Cox-Grimmett coefficient. This condition implies the following moment inequality for sums due to Birkel [4]. If (X_n) is a centered PA sequence with

$$\mathbf{E}|X_1|^{p+\delta} < +\infty, \quad u(n) \leq Bn^{-\frac{(p-2)(p+\delta)}{2\delta}}, \tag{13}$$

for some $p > 2$ and $\delta > 0$, then

$$\mathbf{E}|S_n|^p \leq Kn^{\frac{p}{2}}, \tag{14}$$

where the positive constant K depends only on B, δ and p.

Theorem 7. *Let (X_n) be a centered strictly stationary PA sequence with $0 < \sigma^2 < +\infty$ and such that* (13) *is satisfied. Then the distribution functions F_n of $S_n/(\sigma\sqrt{n})$ converge to Φ in W_p-distance.*
(ii) *If, additionally, c is as in Theorem* 3 (ii), *then* $T_c(F_n, \Phi) \to 0$.

4 Proofs

Proof of Theorem 1. We first show that Π, the set of probability measures on $\mathcal{B}(M \times M)$ with marginals μ and ν, is tight. Indeed, for any positive ε there exist compact sets K_1, $K_2 \in \mathcal{B}(M)$, such that $\mu(K_1) \geq 1 - \frac{\varepsilon}{2}$ and $\nu(K_2) \geq 1 - \frac{\varepsilon}{2}$. Let $\pi \in \Pi$ and let (X, Y) be a random vector with law π. Then,

$$\begin{aligned}\pi(K_1 \times K_2) &= P(X \in K_1, Y \in K_2) = P(X \in K_1) + P(Y \in K_2) - P((X \in K_1) \cup (Y \in K_2)) \\ &\geq \mu(K_1) + \nu(K_2) - 1 \geq (1 - \varepsilon/2) + (1 - \varepsilon/2) - 1 = 1 - \varepsilon.\end{aligned} \tag{15}$$

Since (15) holds for all $\pi \in \Pi$ with the same compact set $K_1 \times K_2$, this proves that Π is tight. Therefore, according to Prokhorov's theorem (Billingsley [3], Section 5), Π is relatively compact.

If $T_c(\mu, \nu) = +\infty$, then $\int c(x, y)d\pi(x, y) = +\infty$, for all $\pi \in \Pi$ and π^* can be chosen to be any probability measure from Π.

If $T_c(\mu, \nu) < +\infty$, then there exists a sequence π_n from Π such that

$$\int c(x, y)d\pi_n(x, y) \to T_c(\mu, \nu). \tag{16}$$

On the other hand, the relative compactness of Π implies the existence of a subsequence π_{n_k} which converges weakly to some probability measure π on $\mathcal{B}(M \times M)$. Let us verify that π is the measure π^* we are looking for. First, we want to prove that $\pi \in \Pi$, i.e. that the marginal distributions of π are μ and ν, respectively.

Let μ_1 and ν_1 be marginals of π. We will check that $\mu_1(B) = \mu(B)$, for any $B \in \mathcal{B}(M)$ such that $\mu_1(\partial B) = 0$. Indeed, since $\partial(B \times M) \subset (\partial B \times M) \cup (B \times \partial M) = \partial B \times M$ (Billingsley [3], (2.8)), we have

$$\pi(\partial(B \times M)) \leq \pi(\partial B \times M) = \mu_1(\partial B) = 0.$$

Therefore, the weak convergence $\pi_{n_k} \Rightarrow \pi$ implies that $\pi_{n_k}(B \times M) \to \pi(B \times M)$, and we obtain

$$\mu(B) = \pi_{n_k}(B \times M) \to \pi(B \times M) = \mu_1(B).$$

Similarly, we can show that $\nu_1(B) = \nu(B)$, for any $B \in \mathcal{B}(M)$ such that $\nu_1(\partial B) = 0$. Finally, it remains to check that two probability measures μ_1 and μ (respectively ν_1 and ν) are the same if they coincide on the Borel sets having a boundary of μ_1-measure (respectively ν_1-measure) zero.

Let $D \in \mathcal{B}(M)$ be a closed set. For $\varepsilon > 0$, let $D^\varepsilon = \{x \in M : d(x, D) < \varepsilon\}$ and let $\mathcal{D} = \{D^\varepsilon, 0 < \varepsilon < 1\}$. Then there exists at most a countable number of ε_k, $0 < \varepsilon_k < 1$, such that sets D^{ε_k} have a boundary of positive μ_1-measure. We remove the sets D^{ε_k} from $\mathcal{D}$, and obtain

$$\mathcal{D}^0 = \{D^\varepsilon, 0 < \varepsilon < 1, \mu_1(\partial D^\varepsilon) = 0\}.$$

We can then choose a decreasing sequence $\varepsilon_n \to 0$, $0 < \varepsilon_n < 1$, with $D_n = D^{\varepsilon_n} \in \mathcal{D}^0$. The sets D_n are such that: (a) $D_{n+1} \subset D_n$ for all n; (b) $\bigcap_n D_n = D \cup \partial D = D$; (c) $\mu_1(D_n) = \mu(D_n)$. The properties (a)–(c) yield

$$\mu_1(D) = \mu_1(\bigcap_n D_n) = \lim_{n\to\infty} \mu_1(D_n) = \lim_{n\to\infty} \mu(D_n) = \mu(D).$$

Therefore, the measures μ_1 and μ coincide on all the closed subsets of M. Since $\mathcal{B}(M)$ is generated by such sets, we conclude that $\mu_1 = \mu$. Similar arguments lead to $\nu_1 = \nu$. We have proved that the probability measure π has respective marginals μ and ν, i.e. $\pi \in \Pi$.

Next, we will check that $\int c(x, y) d\pi(x, y) = T_c(\mu, \nu)$. Since c is lower semicontinuous, for any real b the set $\{(x, y) : c(x, y) > b\}$ is open ([3], Appendix I). Let $A = \{(x, y) : c(x, y) > 0\}$. Then the weak convergence $\pi_{n_k} \Longrightarrow \pi$ and (16) imply that

$$\begin{aligned}
\int c(x, y) d\pi(x, y) &= \int_A c(x, y) d\pi(x, y) \\
&\leq \lim \inf_{n_k} \int_A c(x, y) d\pi_{n_k}(x, y) \\
&= \lim \inf_{n_k} \int c(x, y) d\pi_{n_k}(x, y) \\
&= T_c(\mu, \nu).
\end{aligned}$$

Since $\pi \in \Pi$, the reverse inequality $\int c(x, y) d\pi(x, y) \geq T_c(\mu, \nu)$ holds true. We thus conclude that $\int c(x, y) d\pi(x, y) = T_c(\mu, \nu)$. In other words, the transportation distance becomes the total transportation cost associated to the measure π. Finally, we set $\pi^* = \pi$ and the proof is now complete. □

Proof of Theorem 2 and Corollary 1. Assume that both (a) and (b) are satisfied. Let X, X_n be random elements with respective distributions μ and μ_n and such that X and X_n are independent, for any n. Then $(C(2d(X_n,a)))$ is uniformly bounded, that is $I_1 = \sup_n \mathbf{E}C(2d(X_n,a)) < \infty$. Set $I_2 = \mathbf{E}C(2d(X,a)) < \infty$.

Fix $\varepsilon > 0$ and choose a compact set K_1 in $\mathcal{B}(M)$ such that $\mu(\partial K_1) = 0$ and

$$\int_{K_1^c} C(2d(x,a))d\mu(x) < \varepsilon.$$

The weak convergence $\mu_n \Longrightarrow \mu$ implies the tightness of the family $(\mu_n,\mu)_{n\geq 1}$, thus there exists a compact set $K_2 \in \mathcal{B}(M)$ such that $\mu_n(K_2^c) < \varepsilon$, $\mu(K_2^c) < \varepsilon$ and $\mu(\partial K_2) = 0$. Let $K = K_1 \cup K_2$. Then K is compact, and

$$\int_{K^c} C(2d(x,a))d\mu(x) < \varepsilon, \tag{17}$$

$$\mu_n(K^c) < \varepsilon, \quad \mu(K^c) < \varepsilon, \tag{18}$$

with also $\mu(\partial K) = 0$, since $\mu(\partial K) \leq \mu(\partial K_1) + \mu(\partial K_2)$. Since (b) holds, we can choose a positive integer N_1 such that for any $n \geq N_1$,

$$\left| \int C(2d(x,a))d\mu_n(x) - \int C(2d(x,a))d\mu(x) \right| < \varepsilon. \tag{19}$$

As $X_n \xrightarrow{d} X$, for the chosen compact set K and the continuous function $C(2d(\cdot,a))$ we have

$$\mathbf{E}C(2d(X_n,a))\mathbf{1}_{\{X_n\in K\}} \to \mathbf{E}C(2d(X,a))\mathbf{1}_{\{X\in K\}}.$$

Hence, we can choose a positive integer N_2 such that, for any $n \geq N_2$,

$$\left| \int_K C(2d(x,a))d\mu_n(x) - \int_K C(2d(x,a))d\mu(x) \right| < \varepsilon. \tag{20}$$

Then for $n \geq \max\{N_1, N_2\}$, the estimates (17), (19) and (20) yield

$$\begin{aligned} &\left|\int_{K^c} C(2d(x,a))d\mu_n(x)\right| \\ &\quad \leq \left| \int C(2d(x,a))d\mu_n(x) - \int C(2d(x,a))d\mu(x) \right| \\ &\quad + \left| \int_K C(2d(x,a))d\mu_n(x) - \int_K C(2d(x,a))d\mu(x) \right| + \left| \int_{K^c} C(2d(x,a))d\mu(x) \right| \\ &\quad < 3\varepsilon. \end{aligned} \tag{21}$$

The weak convergence $X_n \xrightarrow{d} X$ implies that $C(2d(X_n,X))\mathbf{1}_{\{X_n\in K,\ X\in K\}} \xrightarrow{d} 0$. The continuous function $C(d(x,y))$ is bounded on the compact set $K \times K$, therefore

$$\mathbf{E}C(d(X_n,X))\mathbf{1}_{\{X_n\in K,\ X\in K\}} \to 0.$$

This means that there exists a positive integer N_3 such that, for any $n \geq N_3$,

$$\left| \int_K \int_K C(d(X_n, X)) d\pi_n(x, y) \right| < \varepsilon, \tag{22}$$

where π_n is the joint distribution of X_n and X.

Since C is a non-negative and non-decreasing,

$$\begin{aligned} C(d(x,y)) &\leq C(d(x,a) + d(y,a)) \leq C(2\max\{d(x,a), d(y,a)\}) \\ &\leq C(2d(x,a)) + C(2d(y,a)), \end{aligned} \tag{23}$$

for all $x, y \in M$.

Using the inequalities (17), (18), (21), (23), and the independence of X_n and X, we have:

$$\begin{aligned} &\int_{K^c} \int_{K^c} C(d(x,y)) d\pi_n(x,y) = \mathbf{E} C(d(X_n, X)) \mathbf{1}_{\{X_n \in K^c,\ X \in K^c\}} \\ &\leq \mathbf{E} C(2d(X_n, a)) \mathbf{1}_{\{X_n \in K^c\}} \mathbf{1}_{\{X \in K^c\}} + \mathbf{E} C(2d(X, a)) \mathbf{1}_{\{X \in K^c\}} \mathbf{1}_{\{X_n \in K^c\}} \\ &\leq \left(\int_{K^c} C(2d(x,a)) d\mu_n(x) \right) \mu(K^c) + \left(\int_{K^c} C(2d(x,a)) d\mu(x) \right) \mu_n(K^c) \\ &< 3\varepsilon^2 + \varepsilon^2, \end{aligned} \tag{24}$$

for all $n \geq \max\{N_1, N_2\}$. Similarly,

$$\begin{aligned} &\int_K \int_{K^c} C(d(x,y)) d\pi_n(x,y) = \mathbf{E} C(d(X_n, X)) \mathbf{1}_{\{X_n \in K,\ X \in K^c\}} \\ &\leq \mathbf{E} C(2d(X_n, a)) \mathbf{1}_{\{X_n \in K\}} \mathbf{1}_{\{X \in K^c\}} + \mathbf{E} C(2d(X, a)) \mathbf{1}_{\{X \in K^c\}} \mathbf{1}_{\{X_n \in K\}} \\ &\leq I_1 \mu(K^c) + \varepsilon \mu_n(K) \\ &< I_1 \varepsilon + \varepsilon, \end{aligned} \tag{25}$$

and

$$\begin{aligned} &\int_{K^c} \int_K C(d(x,y)) d\pi_n(x,y) = \mathbf{E} C(d(X_n, X)) \mathbf{1}_{\{X_n \in K^c,\ X \in K\}} \\ &\leq \mathbf{E} C(2d(X_n, a)) \mathbf{1}_{\{X_n \in K^c\}} \mathbf{1}_{\{X \in K\}} + \mathbf{E} C(2d(X, a)) \mathbf{1}_{\{X \in K\}} \mathbf{1}_{\{X_n \in K^c\}} \\ &\leq 3\varepsilon \mu(K) + I_2 \mu_n(K^c) < 3\varepsilon + I_2 \varepsilon, \end{aligned} \tag{26}$$

for $n \geq \max\{N_1, N_2\}$.

Thus for $n \geq \max\{N_1, N_2, N_3\}$ the inequalities (22), (24)–(26) yield

$$\begin{aligned} T_c(\mu_n, \mu) \leq \mathbf{E} C(d(X_n, X)) &= \int \int C(d(x,y)) d\pi_n(x,y) \\ &= \int_K \int_K C(d(x,y)) d\pi_n(x,y) + \int_{K^c} \int_{K^c} C(d(x,y)) d\pi_n(x,y) \\ &+ \int_K \int_{K^c} C(d(x,y)) d\pi_n(x,y) + \int_{K^c} \int_K C(d(x,y)) d\pi_n(x,y) \\ &\leq \varepsilon(6 + 4\varepsilon + I_1 + I_2). \end{aligned}$$

We conclude that (a) and (b) imply $T_c(\mu_n, \mu) \to 0$.

Next, we assume that $T_c(\mu_n, \mu) \to 0$ and verify that (a) $\mu_n \Longrightarrow \mu$ takes place. According to Theorem 1, for any n there exists a pair of random elements X_n and X with distributions μ_n and μ, respectively, which are minimizers of the total transportation cost: $T_c(\mu_n, \mu) = \mathbf{E}C(d(X_n, X))$. Let us note that X may depend on n, so each time it appears in this proof, we assume that $X = X^{(n)}$. (Of course all the $X^{(n)}$ have the same law μ.)

Since C is a non-negative function, $\mathbf{E}C(d(X_n, X)) \to 0$, that is $C(d(X_n, X)) \xrightarrow{L_1} 0$. This implies that

$$C(d(X_n, X)) \xrightarrow{P} 0. \tag{27}$$

Fix $\varepsilon > 0$. As C is non-decreasing, we have

$$\{d(X_n, X) > \varepsilon\} \subset \{C(d(X_n, X)) \geq C(\varepsilon)\}.$$

The convergence result (27) implies that the probability of the last event tends to 0, for any positive $C(\varepsilon)$, as $n \to \infty$. Hence for any $\varepsilon > 0$, $P(d(X_n, X) > \varepsilon) \to 0$, as $n \to \infty$. The convergence, in probability, of $d(X_n, X) = d(X_n, X^{(n)}) \xrightarrow{P} 0$ implies that $\mu_n \Longrightarrow \mu$ (Billingsley [3], theorem 4.1).

Now, assume that the doubling condition (6) is satisfied and let us verify that $T_c(\mu_n, \mu) \to 0$ implies (b$'$). Since $\mu_n \Longrightarrow \mu$ and since $C(d(\cdot, a))$ is continuous on M, weak convergence holds: $C(d(X_n, a) \xrightarrow{d} C(d(X, a))$. In order to verify (b$'$), it thus suffices to check that the sequence $(C(d(X_n, a)))$ is uniformly integrable. Uniform integrability is equivalent to the pair of conditions: (i) $(\mathbf{E}C(d(X_n, a)))$ is uniformly bounded and (ii) for $A \in \mathcal{F}$, $(\mathbf{E}C(d(X_n, a))\mathbf{1}_A)$ is uniformly continuous, (i.e. $\sup_n \mathbf{E}C(d(X_n, a))\mathbf{1}_A \to 0$ as $P(A) \to 0$).

Together (6) and (23) yield the inequalities

$$C(d(x, a)) \leq \lambda C\left(\frac{1}{2}d(x, a)\right) \leq \lambda C(d(x, y)) + \lambda C(d(y, a)), \tag{28}$$

for all $x, y \in M$ and the positive constant λ. Then

$$\mathbf{E}C(d(X_n, X)) \geq \frac{1}{\lambda}\mathbf{E}C(d(X_n, a)) - \mathbf{E}C(d(X, a)). \tag{29}$$

Suppose that $(\mathbf{E}C(d(X_n, a)))$ is not uniformly bounded. Then there exists a subsequence $(\mathbf{E}C(d(X_{n'}, a)))$ such that $\mathbf{E}C(d(X_{n'}, a)) \to +\infty$. Applying (29) to this subsequence, we come to the following contradiction:

$$0 \leftarrow \mathbf{E}C(d(X_n, X)) \geq \frac{1}{\lambda}\mathbf{E}C(d(X_{n'}, a)) - \mathbf{E}C(d(X, a)) \to +\infty.$$

Thus, $(\mathbf{E}C(d(X_n, a)))$ is uniformly bounded.

Let ε be fixed, and let $A \in \mathcal{F}$. Since $T_c(\mu_n, \mu) \to 0$, we can choose a positive integer N such that $\mathbf{E}C(d(X_n, X))\mathbf{1}_A < \varepsilon$, for all $n \geq N$. By applying once again the inequality (29), we obtain

$$\sup_{n \geq N} \mathbf{E}C(d(X_n, a))\mathbf{1}_A \leq \lambda \sup_n \mathbf{E}C(d(X_n, X))\mathbf{1}_A + \lambda \mathbf{E}C(d(X, a))\mathbf{1}_A.$$

Let $P(A) \to 0$. Since $(C(d(X_n, X)))$ is uniformly integrable and since $\mathbf{E}C(d(X, a)) \leq \mathbf{E}C(2d(X, a)) < \infty$,

$$\sup_{n \geq N} \mathbf{E}C(d(X_n, a))\mathbf{1}_A \to 0,$$

i.e. $(\mathbf{E}C(d(X_n,a))\mathbf{1}_A)$ is uniformly continuous. Hence, the sequence $(C(d(X_n,a)))$ is uniformly integrable and (b′) $\int C(d(x,a))d\mu_n \to \int C(d(x,a))d\mu$ holds.

Note that from (6) and since C is non-decreasing, the following two inequalities hold true

$$C(2d(x,a)) \leq \lambda C(d(x,a)), \qquad C(d(x,a)) \leq C(2d(x,a)),$$

for any $x \in M$. This implies that $(\int C(d(x,a))\mu_n(dx) < \infty) \iff (\int C(2d(x,a))\mu_n(dx) < \infty)$ and that $(\int C(d(x,a))\mu(dx) < \infty) \iff (\int C(2d(x,a))\mu(dx) < \infty)$. Therefore, in the setting of the theorem, the sequences $(C(d(X_n,a)))$ and $(C(2d(X_n,a)))$ are both either uniformly integrable or not, and (b)$\iff$ (b′).

This observation completes the proof of Theorem 2 and of Corollary 1. □

Proof of Corollary 2. Let K_1 and K_2 be the respective supports of μ and ν. If μ and ν are absolutely continuous with respective densities f_1 and f_2, then

$$\begin{aligned}\left|\int \phi(x)d\mu - \int \phi(x)d\nu\right| &= \left|\int \phi(x)f_1(x)dx - \int \phi(x)f_2(x)dx\right| \\ &\leq \int_{K_1\cup K_2} |\phi(x)||f_1(x) - f_2(x)|dx \\ &\leq L_\phi \|\mu - \nu\|_{TV},\end{aligned}$$

where $L_\phi = \sup\{|\phi(x)| : x \in \overline{(K_1 \cup K_2)}\}$ (here $\overline{A} = A \cup \partial A$).

To prove the result in the general case, let the partition $(A_m)_{m\in\mathbf{Z}}$ of M, $A_m \in \mathcal{B}(M)$, be defined as follows:

$$A_m = \{x \in M : m - 1 \leq \phi(x) < m\}.$$

Thus,

$$\begin{aligned}\left|\int \phi(x)d\mu - \int \phi(x)d\nu\right| &= \left|\sum_{m=-\infty}^{+\infty} \int \phi(x)\mathbf{1}_{A_m} d(\mu - \nu)\right| \\ &\leq \sum_{m=-\infty}^{+\infty} |m||\mu(A_m) - \nu(A_m)| \\ &\leq L_\phi \|\mu - \nu\|_{TV},\end{aligned} \tag{30}$$

where L_ϕ is defined as above, and where we used the dual definition of the total variation distance.

Let μ_n and μ be probability measures on M with compact supports respectively denoted K_n and K. Let also $\cup_n K_n$ be bounded and $\|\mu_n - \mu\|_{TV} \to 0$. Convergence in total variation implies weak convergence $\mu_n \Longrightarrow \mu$. All the conditions of Theorem 2 are satisfied. Therefore, to prove the convergence of μ_n to μ in T_c, it suffices to check that $\int C(2d(x,a))d\mu_n \to \int C(2d(x,a))d\mu$, for some $a \in M$. The inequality (30) yields, for any n,

$$\left|\int \phi(x)d\mu_n - \int \phi(x)d\mu\right| \leq L_\phi \|\mu_n - \mu\|_{TV}, \tag{31}$$

where $L_\phi = \sup\{|\phi(x)| : x \in \overline{\cup K_n}\} < \infty$ does not depend on n. By fixing $a \in M$ and applying (31) to $\phi(x) = C(2d(x,a))$, we obtain that convergence in total variation implies the convergence of the integrals $\int \phi(x)d\mu_n \to \int \phi(x)d\mu$. This completes the proof. □

Proof of Theorem 3. (i) First, we prove the statement for $p > 2$. Let F be the distribution function of $X_1, X_2, \ldots$ and, for a fixed n let $\eta_1, \eta_2, \ldots, \eta_n$ be a set of i.i.d. normal random variables with mean zero and variance $1/n$. According to a result of Sakhanenko ([19], Theorem 5), there exist i.i.d. random variables $\xi_1, \xi_2, \ldots, \xi_n$, each having the distribution function $F_{\xi_i}(x) = F(x\sigma\sqrt{n})$ and such that

$$\mathbf{E} \max_{1\le k\le n} \left| \sum_{i=1}^{k} \xi_i - \sum_{i=1}^{k} \eta_i \right|^p \le Cp^{2p} \sum_{i=1}^{n} \mathbf{E}|X_i|^p. \tag{32}$$

Therefore, for $Z_n = \sum_{i=1}^n \eta_i$, $Z_n \sim N(0,1)$, the following chain of inequalities is true:

$$\begin{aligned} W_p^p(F_n, \Phi) &\le \mathbf{E}\left| \frac{S_n}{\sigma\sqrt{n}} - Z_n \right|^p \le \mathbf{E}\left| \sum_{i=1}^{n} \xi_i - \sum_{i=1}^{n} \eta_i \right|^p \\ &\le Cp^{2p} n \frac{\mathbf{E}|X_1|^p}{\sigma^p n^{\frac{p}{2}}} = Cp^{2p} n^{1-\frac{p}{2}} \sigma^{-p} \mathbf{E}|X_i|^p \to 0, \end{aligned} \tag{33}$$

as $n \to \infty$.

Now, let $p = 2$. Let $Z \sim N(0,1)$ and the sequence (X_n) be independent. We set $Y_n = S_n/(\sigma\sqrt{n})$ and verify that $\mathbf{E}Y_n^2 \to \mathbf{E}Z^2$.

Fix $\varepsilon > 0$ and choose a compact set K, $K \in \mathcal{B}(\mathbf{R})$, such that

$$\gamma(K^c) < \varepsilon, \quad \mathbf{E}Z^2 \mathbf{1}_{\{Z\in K^c\}} < \varepsilon, \quad \gamma(\partial K) = 0, \tag{34}$$

where γ is the standard Gaussian measure on $\mathbf{R}$. Using (34) and the convergence of $\mathbf{E}Y_n^2 \mathbf{1}_{\{Z\in K\}}$ to $\mathbf{E}Z^2 \mathbf{1}_{\{Z\in K\}}$, we obtain

$$\left|\mathbf{E}Y_n^2 - \mathbf{E}Z^2\right| \le \left|\mathbf{E}Y_n^2 \mathbf{1}_{\{Z\in K\}} - \mathbf{E}Z^2 \mathbf{1}_{\{Z\in K\}}\right| + \mathbf{E}Y_n^2 \mathbf{1}_{\{Z\in K^c\}} + \mathbf{E}Z^2 \mathbf{1}_{\{Z\in K^c\}} \le 3\varepsilon$$

for sufficiently large n. Therefore, $\mathbf{E}Y_n^2 \to \mathbf{E}Z^2$.

Finally, Corollary 1 gives $W_2(F_n, \Phi) \to 0$. This completes the proof of (i).

(ii) The convergence result $W_p(F_n, \Phi) \to 0$ and the theorem of Bickel and Freedman together imply that $\mathbf{E}|Y_n|^p \to \mathbf{E}|Z|^p$. We will show that for the given function C all the conditions of Theorem 2 are satisfied. Indeed, we have $\mathbf{E}C(2|Z|) < +\infty$, while the finiteness of $\int C(2|x|)dF_n(x)$ follows from the convergence of the absolute p-moments combined with

$$C(2|Y_n|) = C(2|Y_n|)\mathbf{1}_{\{|Y_n|\le x_0\}} + C(2|Y_n|)\mathbf{1}_{\{|Y_n|>x_0\}} \le C(x_0)\mathbf{1}_{\{|Y_n|\le x_0\}} + \beta|Y_n|^p \mathbf{1}_{\{|Y_n|>x_0\}}, \tag{35}$$

with a positive constant β such that $C(x) \le \beta x^p$, for all $x > x_0$. The CLT provides the weak convergence result $F_n \Longrightarrow \Phi$, while we obtain $\mathbf{E}C(2|Y_n|) \to \mathbf{E}C(2|Z|)$ from the uniform integrability of $(C(2|Y_n|)$ which follows from (35) and the uniform integrability of $(|Y_n|^p)$. Thus, applying Theorem 2, we obtain that $T_c(\mu_n, \gamma) \to 0$. This concludes the proof of Theorem 3. □

Proof of Theorem 4. (i) This closely follows arguments of the proof of Theorem 3 (i). Once again we apply Sakhanenko's inequality (32), this time for independent variables ξ_1, ..., ξ_n with distribution functions $F_{\xi_i}(x) = F_{X_i}(x\sigma_n)$, $i = 1, \ldots, n$. As in getting (33), we have

$$W_p^p(F_n, \Phi) \le Cp^{2p} \frac{\sum_{i=1}^n \mathbf{E}|X_i|^p}{\sigma_n^p} \to 0,$$

as $n \to \infty$, using the Lyapunov condition (7).
The proofs of part (ii) of Theorem 4 to 7 with obvious changes repeat the proof of Theorem 3 (ii). □

Proof of Theorem 5. (i) We will check that all the conditions of Corollary 1 are satisfied. Doukhan *et al.* showed that if $\mathbf{E}|X_1|^k < +\infty$ and $\sum_{n=1}^{\infty} n^{1/(k-1)}\alpha_n < \infty$ for $k > 2$, then the quantile condition (8) holds. It is easy to verify that the series above converges for $k = p + \delta$ thanks to the rate condition (9); thus, the α-mixing CLT is valid and $F_n \Longrightarrow \Phi$.

Next, we verify that $\mathbf{E}|Y_n|^p \to \mathbf{E}|Z|^p$, where $Y_n = S_n/(\sigma\sqrt{n})$. The CLT gives $Y_n \xrightarrow{d} Z$, while the uniform boundedness of $\mathbf{E}|Y_n|^p$ follows from the moment bound (10):

$$\mathbf{E}|Y_n|^p \leq K\sigma^{-p}(\mathbf{E}|X_1|^{p+\delta})^{\frac{p}{p+\delta}}. \tag{36}$$

Let $Z \sim N(0,1)$ and the sequence (X_n) be independent. Fix $\varepsilon > 0$ and choose a compact set K as in (34) (with 2 replaced by p). From (36) and the convergence of $\mathbf{E}|Y_n|^p\mathbf{1}_{\{Z\in K\}}$ to $\mathbf{E}|Z|^p\mathbf{1}_{\{Z\in K\}}$, we get

$$\begin{aligned}|\mathbf{E}|Y_n|^p - \mathbf{E}|Z|^p| &\leq \left|\mathbf{E}|Y_n|^p\mathbf{1}_{\{Z\in K\}} - \mathbf{E}|Z|^p\mathbf{1}_{\{Z\in K\}}\right| + \mathbf{E}|Y_n|^p\mathbf{1}_{\{Z\in K^c\}} + \mathbf{E}|Z|^p\mathbf{1}_{\{Z\in K^c\}} \\ &\leq \varepsilon + \varepsilon K\sigma^{-p}(\mathbf{E}|X_1|^{p+\delta})^{\frac{p}{p+\delta}} + \varepsilon,\end{aligned}$$

for sufficiently large n. Therefore, $\mathbf{E}|Y_n|^p \to \mathbf{E}|Z|^p$.

Finally, Corollary 1 gives $W_p(F_n, \Phi) \to 0$. □

Proofs of Theorems 6 (i) and 7 (i). These are carried out by using the same arguments as in the proof of Theorem 5 (i). Instead of the moment bound (10), we use the condition (12) and the Rosenthal inequality for independent sequence (X_n^*) to prove Theorem 6 (i) and Birkel's result (14) to prove Theorem 7 (i).

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