

No books or notes allowed. No laptop or wireless devices allowed. Show all your work for full credit. **Write clearly and legibly.**

Name (print): _____

Question:	1	2	3	4	5	Total
Points:	40	20	30	20	0	110
Score:						

Question:	1	2	3	4	5	Total
Bonus Points:	0	0	0	0	20	20
Score:						

Question 1 40 point

Let X be a continuous r.v. with p.d.f. given by

$$f(x) = \begin{cases} p + 2(1-p)x & 0 < x < 1 \\ 0 & \text{otherwise} \end{cases}$$

where p is a parameter.

- (a) (10 points) for which value of p is f a valid p.d.f. (**Hint:** remember that there are 2 conditions you should check.)

Solution: We have

$$\int_0^1 f(x)dx = 1$$

for every p so that we just need to check that $f(x) > 0$. Since f is linear in x it is enough that $f(0) > 0$ and $f(1) > 0$. Thus we get

$$0 \leq p \leq 2.$$

- (b) (15 points) Compute the expected value $\mathbb{E}(X)$ and the variance $V(X)$ of X .

Solution: We have

$$\mathbb{E}(X) = \int_0^1 xf(x)dx = p \int_0^1 xdx + 2(1-p) \int_0^1 x^2dx = \frac{p}{2} + \frac{2}{3}(1-p) = \frac{2}{3} - \frac{p}{6}$$

and

$$\mathbb{E}(X^2) = \int_0^1 x^2f(x)dx = p \int_0^1 x^2dx + 2(1-p) \int_0^1 x^3dx = \frac{p}{3} + \frac{1}{2}(1-p) = \frac{1}{2} - \frac{p}{6}$$

so that

$$V(x) = \frac{1}{36} (p(2-p) + 2)$$

(c) (15 points) Show that

$$\mathbb{E} \left(\left(X - \frac{2}{3} \right)^2 \right) \geq \left(\frac{p}{6} \right)^2$$

(**Hint:** use Jensen inequality.)

Solution: Use Jensen inequality to get

$$\mathbb{E} \left(\left(X - \frac{2}{3} \right)^2 \right) \geq \left(\mathbb{E}(X) - \frac{2}{3} \right)^2 = \left(\frac{1}{2} - \frac{p}{6} - \frac{1}{2} \right)^2 .$$

Question 2 20 point

Let $N_k, k = 1, 2, 3, \dots$, be an infinite sequence of geometric random variable with parameter $p_k = \frac{\lambda}{k}$, that is

$$\mathbb{P}(N_k = n) = (1 - p_k)^{n-1} p_k \quad \text{for } n \geq 1.$$

and $\mathbb{P}(N_k = n) = 0$ for $n < 1$. Moreover let Y be an exponential r.v. with parameter λ , that is

$$f_Y(y) = \lambda e^{-\lambda y} \quad \text{for } y \geq 0$$

and $f_Y(y) = 0$ for $y < 0$.

Show that $Z_k = N_k/k$ converge in distribution to Y as $k \rightarrow \infty$. (**Hint:** compute the c.d.f. of Z_k , that is $F_k(x) = \mathbb{P}(Z_k \leq x)$ for every real number x .)

Solution: Observe that we have

$$\mathbb{P}(Z_k \leq x) = \mathbb{P}(N_k \leq kx) = \sum_{i=1}^{\lfloor kx \rfloor} (1 - p_k)^{i-1} p_k = 1 - (1 - p_k)^{\lfloor kx \rfloor - 1} = \left(1 - \frac{\lambda}{k}\right)^{\lfloor kx \rfloor - 1}$$

while

$$\mathbb{P}(Y \leq y) = 1 - e^{-\lambda y}.$$

Observe that

$$\frac{\lfloor kx \rfloor - 1}{k} \rightarrow x$$

as $k \rightarrow \infty$ so that

$$\left(1 - \frac{\lambda}{k}\right)^{\lfloor kx \rfloor - 1} = \left(\left(1 - \frac{\lambda}{k}\right)^k\right)^{\frac{\lfloor kx \rfloor - 1}{k}} \xrightarrow{k \rightarrow \infty} e^{-\lambda x}$$

for every x .

Question 3 30 point

A student is attempting a multiple choices exam. For each question there are 4 possible answers. He has a probability of 0.75 of knowing the correct answer. If he does not know the answer he chooses one answer uniformly and randomly. All questions and answers are independent.

To get a B he need to answer correctly 85% of the questions while to get an A he needs to answer correctly 95% of the questions.

- (a) (15 points) If the test contains 40 questions, use a normal approximation (CLT) and the table provided to compute the probability p_B that the student will get at least a B and the probability p_A that the student will get a A.

Solution: Let p be the probability that the student give a correct answer. We have

$$p = 0.75 + 0.25 \cdot 0.25 = 0.8125.$$

Let X_i be 1 if he answer correctly to the i -th question and 0 otherwise. Thus $\mathbb{E}(X_i) = 0.8125$ and $V(X_i) = 0.1523$. We get

$$\begin{aligned} p_B &= \mathbb{P}\left(\sum_{i=1}^{40} X_i > 0.85 \cdot 40\right) = \\ &= \mathbb{P}\left(\frac{\sum_{i=1}^{40} X_i - 0.8125 \cdot 40}{0.390\sqrt{40}} > \frac{(0.85 - 0.8125) \cdot 40}{0.390\sqrt{40}}\right) = \\ &= 1 - \Phi(0.60) = 0.274 \end{aligned}$$

while

$$\begin{aligned} p_B &= \mathbb{P}\left(\sum_{i=1}^{40} X_i > 0.95 \cdot 40\right) = \\ &= \mathbb{P}\left(\frac{\sum_{i=1}^{40} X_i - 0.8125 \cdot 40}{0.390\sqrt{40}} > \frac{(0.95 - 0.8125) \cdot 40}{0.390\sqrt{40}}\right) = \\ &= 1 - \Phi(2.23) = 0.013. \end{aligned}$$

- (b) (15 points) Let p_B be the probability that a student that knows 75% of the answers will get a B or more. If the teacher wants p_B to be less than 0.025, how many question should there be on the exam.

Solution: He wants to find N such that

$$\mathbb{P}\left(\sum_{i=1}^N X_i > 0.85 \cdot N\right) \leq 0.025$$

This means

$$\mathbb{P}\left(\frac{\sum_{i=1}^N X_i - 0.8125 \cdot N}{0.390\sqrt{N}} > \frac{(0.85 - 0.8125) \cdot \sqrt{N}}{0.390}\right) = 1 - \Phi(0.096 \cdot \sqrt{N}) \leq 0.025$$

From the table we

$$\Phi(1.96) = 0.975$$

so that he needs

$$N > \left(\frac{1.96}{0.096}\right)^2 = 416$$

questions.

Question 4 20 point

Let Z be a standard normal random variable. Find the n -th moment $m_n = \mathbb{E}(Z^n)$ of Z , for every n . (**Hint:** you can use integration by part to relate m_n with m_{n-2} and then use induction. Alternatively you can use the Taylor expansion around 0 of the m.g.f of a normal standard.)

Solution: By symmetry we just need to look at even n , that is $n = 2k$.

First method: integrating by part we get

$$\begin{aligned} m_{2k} &= \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} z^{2k} e^{-\frac{z^2}{2}} dz = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} z^{2k-1} z e^{-\frac{z^2}{2}} dz = \\ &= (2k-1) \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} z^{2k-2} e^{-\frac{z^2}{2}} dz = (2k-1)m_{2k-2} \end{aligned}$$

we know that $m_0 = 1$ so that

$$m_{2k} = (2k-1)(2k-3)(2k-5) \cdots 5 \cdot 3 \cdot 1$$

Second method: We know that the moment generating function of Z is

$$M_Z(t) = e^{\frac{t^2}{2}} = \sum_{k=0}^{\infty} \frac{t^{2k}}{2^k k!}$$

so that

$$\frac{d^k}{dt^k} M_Z(t) \Big|_{t=0} = \frac{(2k)!}{2^k k!} = (2k-1)(2k-3)(2k-5) \cdots 5 \cdot 3 \cdot 1$$

5. (20 points (bonus)) Let N_1 , N_2 and N_3 be discrete random variables with joint probability mass function

$$p(n_1, n_2, n_3) = \mathbb{P}(N_1 = n_1 \& N_2 = n_2 \& N_3 = n_3) = \frac{3^{-N} N!}{n_1! n_2! n_3!}$$

if $n_1 + n_2 + n_3 = N$ and 0 otherwise.

Compute the marginal mass function p_{N_1} of N_1 , that is

$$p_{N_1}(n_1) = \mathbb{P}(N_1 = n_1)$$

and the conditional mass function $p_{N_2, N_3 | N_1}$ of N_2 and N_3 given N_1 , that is

$$p_{N_2, N_3 | N_1}(n_2, n_3 | n_1) = \mathbb{P}(N_2 = n_2 \& N_3 = n_3 | N_1 = n_1).$$

(**Hint:** you can answer the question without doing any computation. Think what situation is described by N_1 , N_2 and N_3 .)

Solution: Observe that N_1 , N_2 and N_3 are the result of repeating an experiment with 3 possible equiprobable outcomes (say 1,2,3) N times. $\mathbb{P}(N_1 = n_1)$ represents the probability of obtaining n_1 1s when the probability of a 1 is $1/3$. Thus $\mathbb{P}(N_1 = n_1)$ is a binomial with $p = 1/3$ that is

$$\mathbb{P}(N_1 = n_1) = \frac{N!}{(N - n_1)! n_1!} \left(\frac{1}{3}\right)^{n_1} \left(\frac{2}{3}\right)^{N - n_1}.$$

On the other hand if you know you had exactly n_1 1's, the remaining outcomes are 2 or 3, with equal probability. Thus

$$p_{N_2, N_3 | N_1}(n_2, n_3 | n_1) = \frac{2^{-(N - n_1)} (N - n_1)!}{n_2! n_3!}$$

if $n_2 + n_3 = N - n_1$ and 0 otherwise.

In formulas we have

$$\begin{aligned} p_{N_1}(n_1) &= \sum_{n_2, n_3} p(n_1, n_2, n_3) = \sum_{n_2 + n_3 = N - n_1} \frac{3^{-N} N!}{n_1! n_2! n_3!} = \\ &= \frac{3^{-N} 2^{N - n_1} N!}{(N - n_1)! n_1!} \sum_{n_2 + n_3 = N - n_1} \frac{2^{-(N - n_1)} (N - n_1)!}{n_2! n_3!} = \\ &= \frac{N!}{(N - n_1)! n_1!} \left(\frac{1}{3}\right)^{n_1} \left(\frac{2}{3}\right)^{N - n_1} \end{aligned}$$

so that N_1 is a binomial r.v. with N trials and $p = 1/3$.

Moreover we have

$$\begin{aligned} p_{N_2, N_3 | N_1}(n_2, n_3 | n_1) &= \frac{3^{-N} N!}{n_1! n_2! n_3!} \left(\frac{N!}{(N - n_1)! n_1!} \left(\frac{1}{3}\right)^{n_1} \left(\frac{2}{3}\right)^{N - n_1} \right)^{-1} = \\ &= \frac{2^{-(N - n_1)} (N - n_1)!}{n_2! n_3!} \end{aligned}$$

if $n_2 + n_3 = N - n_1$ and 0 otherwise.

Thus N_2 is a binomial r.v with $N - n_1$ trials and $p = 1/2$.

Useful Formulas

- **Normal Distribution:** if Z is a standard normal r.v. then its density function is

$$f(z) = \frac{1}{\sqrt{2\pi}} e^{-\frac{z^2}{2}}$$

while $E(Z) = 0$ and $V(Z) = 1$. The moment generating function $M_Z(t)$ is given by

$$M_Z(t) = e^{\frac{t^2}{2}}.$$

Moreover

$$\Phi(z) = \mathbb{P}(Z \leq z)$$

is given in the table on next page. Finally if X is normal with $\mathbb{E}(X) = \mu$ and $V(X) = \sigma^2$ then

$$Y = \frac{X - \mu}{\sigma}$$

is normal standard.

- **Jensen's Inequality:** If X is a r.v. and g is a convex function then

$$\mathbb{E}(g(X)) \geq g(\mathbb{E}(X)).$$

- **CLT:** if X_i is a sequence of i.i.d. random variable with expected value μ and variance σ^2 and

$$S_n = \frac{1}{\sqrt{n}} \sum_{i=1}^n \frac{X_i - \mu}{\sigma}$$

then S_n converges in distribution to a normal standard r.v. Z .

- **Convergence in Distribution:** we say that the sequence X_n converge in distribution to X if

$$\mathbb{P}(X_n \leq x) \rightarrow_{n \rightarrow \infty} \mathbb{P}(X \leq x)$$

for every $x \in \mathbb{R}$.

